Starving and Deceiving? How Disasters Reshape Politicians’ Incentives to Lie

Shuo Chen, Xinyu Fan, and Xuanyi Wang

Abstract

How do disasters reshape politicians’ incentives to lie? Using a natural experiment during 1959-1961 in China, the paper empirically identifies and tests the impacts of famine exposure on local politicians’ incentives to lie about economic performances. We measure such lies by quantifying GDP manipulation through machine learning XGBoost techniques, then use a difference-in-differences approach to show that leaders who experienced famine in childhood are less likely to lie. To identify the underlying mechanism, we utilize the World Values Survey (WVS) data to reveal that disasters make famine-exposed individuals more risk-averse, thus more reluctant to lie.

Keywords: Famine, Lies, Politicians, GDP Manipulation

JEL: D91, H83, J18
The honesty of leaders is appreciated as a crucial virtue—across time, institutions, and civilizations, from Cicero (“Where is there dignity unless there is honesty?”) to Edward Sandys (“Honesty is the best policy.”). Conversely, if leaders lie, these lies can be highly detrimental to public interests: they can disrupt the formation of public expectations and convey incorrect information that the public is forced to rely upon (Callander and Wilkie, 2007; Marro, 1985). Lying to the public is a risky maneuver for a politician because the discovery of the lie may lead to social unrest and even to civil strife. Consequently, knowing why politicians lie would help to better protect public interests. Existing literature suggests that exposure to disaster can drastically reshape behaviors at individual levels (Brooke et al., 2015; Lonigan et al., 1991; Shore et al., 1986; Yuriko et al., 2017). However, it remains largely unknown whether the exposure to disaster affects public servants’ propensity to lie. Such knowledge is veiled by two challenges: First, lies are hard to quantify and to analyze empirically. Second, establishing the causality between disasters and lying requires clean identification. This paper confronts these two challenges in context of China by means of the country’s unique institutional organization structure, where the measurement problem of lies is solved by estimating the fraudulent GDP manipulation of local officials, and the identification problem is solved by referring to a natural experiment, namely, the great famine in China (1959-1961).

The great famine induced by the Great Leap Forward campaign in China caused more than 17 million lost lives nationwide during 1959-1961, marking the tragedy as one of the biggest famines of human history (Kung and Chen, 2011). Some of the children who suffered in the famine but survived to serve as county-level leaders in later years. We collect their personal information, among 5,383 county-level leaders in 2,161 counties in China, from 2000-2013, which provides a unique opportunity to cross-compare how their behavior in public offices were influenced by their childhood famine exposures. Meanwhile, the top-down personnel control in China promotes public officials based upon an evaluation of
local economic performance (Li and Zhou, 2005). Substantial career incentives thus stimulate local officials to inflate local GDP figures either through developing the economy, or through manipulating the performance data (Charbonneau and Bellavance, 2012). To quantify politicians’ lies, we use novel XGBoost machine learning techniques to recover the true values of local GDP from a series of less manipulable indices, such as night light intensities and air pollution, and compare with the reported values to determine the degree of manipulation. By linking public officials’ identities with the severity of their exposure to famine in childhood, we use a difference-in-differences approach to investigate the effect of exposure on their predisposition to lie in reporting GDP figures.

Our reduced form result shows that public officials with childhood (0-6 years) exposure to the great famine of 1959-1961 are 4.2% less manipulative in local GDP calculations, and the reduction in manipulation is more significant for public officials whose hometowns were more heavily struck by the disaster. Moreover, the effect is more substantial for officials who were younger during the great famine, with the magnitude shrinking as the age grows. This is consistent with the existing literature that early year experiences have more profound long-term impacts (Adhvaryu et al., 2018; 2019; Maccini and Yang, 2009; Malmendier and Nagel, 2011; Sviatschi, 2018). We further conduct a series of robustness checks to validate the baseline results and show that famine exposures reshape public servants’ attitudes toward lying, instead of selecting a few conservative or honest individuals.

Next, we explore the mechanism that restrains GDP manipulation. Psychological and medical literature point to the exposure to negative experience—famine, in our context—that leads individuals to become increasingly risk-averse. Given the high stake of GDP manipulation, we conjecture that famine exposures altered officials’ risk preferences, which in turn made them more reluctant to lie. Since the officials’ risk preferences are not readily available, we utilize the World Values Survey (WVS) data to test a similar group of famine-
exposed individuals and show that they are indeed more risk-averse than the non-exposed, even as they share similar personal traits in other social preferences such as altruism, trust, social network, happiness, and time preferences. This suggests that the changes in risk preferences may be the reason for restraint on the motivation to lie.

This paper deepens our understanding of the long-term impacts of disasters. Recent literature has studied the change of physical and mental conditions of individuals after tsunamis, hurricanes, wars, and financial crises (Almond et al., 2010; Chen and Zhou, 2007; Currie and Vogl, 2013; Kesternich et al., 2014); of labor market performance and human capital (Alderman et al., 2001; Belasen and Polachek, 2008; Currie et al., 2010; Ewing et al., 2009; Handa and Peterman, 2007; Oreopoulos et al., 2008); and of risk preferences (Becchetti et al., 2017; Brown et al., 2018; Callen, 2015; Eckel et al., 2009; Guiso et al., 2018; Kim and Lee, 2014). Our work contributes to this strand of research in two fronts. First, compared to many of the long-term shocks reported in the literature, the famine exposure lasted for three years, which enables us to capture better the changes in individual behavior patterns and underlying preferences. Second, most existing studies focus on the behavior of individuals, regardless of their occupational externality. In this paper, we focus on a homogenous group with large social externalities—the public servants. We find that childhood exposure to disasters has persistent effects in shaping the risk preferences and, moreover, the public servants’ policy preferences. Furthermore, this paper provides an additional channel of disasters’ influence on economic development, namely, that disasters reshape economic

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2 For instance, Cassar et al. (2017) employ the Indian Ocean tsunami in 2004 as a natural experiment; Hanaoka et al. (2018) study the Great East Japan Earthquake in 2008. The disasters in these studies lasted for a few days, and the impacts were restricted to specific regions within a country. As comparison, the great famine lasted for three years, and had profound impact all across China, which allows us to study regional and temporal variations.

3 The sample of officials features another unique advantage, compared to individual samples: the political selection process in China is highly meritocratic, thus the selected officials are often homogenous in competence and cognitive abilities. Therefore, omitted variable bias is less of a concern.

4 Existing channels include: 1. Natural disasters resulted in direct political and institutional changes, which consequently led to economic development (Bello et al., 2016; Cavallo et al., 2003); 2. Natural disasters are perceived as punishments from the God, thus triggering socioeconomic changes in theocracies (Skidmore and Toya, 2002).
development by altering the risk preferences and, therefore, the policy preferences of future policymakers.

More broadly, we locate the paper in the literature of political selection. Existing explanations of bureaucratic behavior are confined largely to the discussions of institutions, such as elections in democracies (Alesina and Rodrik, 1994; Besley and Case, 1995; Besley et al., 2010; Ferraz and Finan, 2011; Levitt, 1997) or promotions in non-democracies (Li and Zhou, 2005; Lü and Landry, 2018). Recently, research has begun to focus on the impact of individual characteristics on political actions, such as gender (Ban and Rao, 2008; Chattopadhyay and Duflo, 2004), education and human capital (Besley et al., 2005; He and Wang, 2017), and competence (Chen et al., 2005; Xi et al., 2018; Yao and Zhang, 2015). This paper delves into the discussion by exploring how varied individual risk preferences may serve as a novel perspective to explain bureaucrats’ actions.

I. Background

A. The Great Famine, 1959-1961

The 1959-1961 famine in China was one of the most severe famines in human history (Kung and Chen, 2011). Scholars estimate total mortality at 17 million to 45 million over these three years. Chronologically, the famine started in 1959, with annualized mortality rates of 14.60‰, which was 3.10‰ more than the average mortality rate during 1955-57 (11.49‰). The mortality rate peaked at

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5 Similar with political selection, the management literature shares the emphasis of personal features of leaders. Benmelech and Frydman (2015), Bernile et al. (2017), Chen et al. (2021), Cronqvist and Yu (2017) and Sunder et al. (2017) investigate CEO experience in the military, piloting planes, raising daughters, and surviving disasters, respectively, in order to reveal the impact of personal experience on corporate behaviors. All the results confirm the significance of personal experience, especially early-year experience on behaviors. It thus reiterates the urge to conduct similar analysis in political economy setups.

6 Jiang and Li (1988, 1996) estimated abnormal death of 17 million. Due to the governmental background of the authors, this was commonly regarded as a lower bound of the population loss. Other notable estimations include: 23 million from Peng (1987), 27 million from Coale (1981), 29.5 million from Ashton et al. (1984), 32.46 million from Cao (2005a; b), and 45 million from Dikotter (2010).
more than 25‰, and started to decline since 1961 and returned to normal beginning in 1962. Panel A in Figure 1 depicts the mortality rate dynamics in China from 1955 to 1966. Panel B, in addition, depicts the regional variation of famine severity between 1959 and 1961: the most damaged regions were mid-China provinces. Anhui, one of the most severely hit regions, had a mortality rate of 56.87‰ in 1960, whereas northeastern provinces were less affected: The mortality rate of Liaoning province in the same year amounted to 2.23‰, which was only 4% of Anhui’s. The significant regional variation enables us to investigate the impacts of famine exposures empirically.

![Figure 1. Impacts of the famine, 1955-1966.](image)

**Panel A:** Mortality Rates; **Panel B:** Regional Distribution

Source: Department of Comprehensive Statistics, National Bureau of Statistics of China (1990)

**Figure 1.** Impacts of the famine, 1955-1966.

**B. China’s GDP Reliability and Manipulation**

There have been longstanding doubts concerning of statistics credibility in China, from environmental protection (Andrews, 2008; Chen et al., 2012; Ghanem and Zhang, 2014), and mining accidents (Fisman and Wang, 2017), to pricing indices (Young, 2003). Given the size of the Chinese economy, it is no surprise that the spotlight is on the credibility of the GDP accounts (Chen et al., 2019; Clark et al., 2019; 2002; Holtz, 2014; Morris and Zhang, 2018; Rawski, 2001). The institutional design of the statistics system in China leaves room for GDP manipulations. The
local statistical bureau is *de jure* a subsidiary of the National Bureau of Statistics, but is under *de facto* supervision of the local government. Moreover, the personnel decisions of the local statistical bureaus (e.g., evaluations and promotions) and their budgets are determined by local government leaders (Chen et al., 2020a; Chen et al., 2020b). Meanwhile, all published numbers are subject to approval by local leaders, especially by party secretaries (Pan and An, 2003). Given the promising career prospects and the loopholes in the institutional design, local leaders have both the incentives and the capacity to manipulate GDP reports.

Existing research has explored various approaches to confirm bias in GDP reports. For instance, one such approach has observed discontinuity in terms of the distribution of GDP growth rate reports around growth targets and jump-ups where reported figures surpass their targets. We depict such discontinuity in the online appendix Figure B1 and further show that such discontinuity disappears if we manually increase or decrease GDP targets by 0.1%. Incidents of GDP manipulation are also revealed in reality. In 2015, a county party secretary forced a local private enterprise to over-report revenue from 200 million RMB to 800 million RMB (Xinhua Net, 2015). Furthermore, in early 2018, the provinces of Liaoning, Inner Mongolia, and Tianjin admitted over-reporting GDP in past years by 25% to 33%, respectively (People’s Daily, 2018b; Xinhua Net, 2018). The severity and prevalence of GDP manipulation thus provide a unique opportunity to develop a metric of political lies.

II. Data

A. Data of County Party Secretaries

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8 The discontinuity in GDP is not unique. Chen et al. (2012) and Ghanem and Zhang (2014) find that once Air Pollution Index (API) are included in the evaluation, similar discontinuity also features around the target, which suggests the wide existence of data manipulation.

9 An alternative explanation of such discontinuity points to objective management, as often seen in corporations, especially for revenue management. However, such micro-management techniques are often overly complicated in governments, where GDP data consist of various industries and numerous firms in each industry.
We collect the data of county party secretaries from yearbooks and Baidu Baike, the largest search engine in China and one of the most widely used sources for official information. Specifically, we first collect the names of the party secretaries in office during 2000-2013 from provincial and prefectural yearbooks and then search for their vitae in Baidu Baike to eventually construct the panel that covers 5,383 county party secretaries in 2,161 counties from 2000 to 2013.

Our main independent variable is the childhood exposure to famine for the county party secretaries. To construct the variable, we first classify the age range to identify whether they experienced the famine in their childhood. Specifically, we define 0-6 years as the childhood of an individual, which is a critical stage for child development (Belsky et al., 1991; Nelson, 1993; Stern, 2017). Next, we define childhood famine exposure as whether an individual’s childhood intersected with the 1959-1961 famine period. We further refer to regional-specific famine excess mortality rates to approximate the severity of the famine. The excess mortality rate is calculated in two steps: First, based on birth cohort information acquired from the 1990 census data, we use non-famine cohorts to interpolate counterfactual birth cohort sizes of 1959-1961; second, we calculate the difference between the factual cohort sizes and the interpolated sizes, thus the excess mortality rate (per Meng et al., 2015).

B. GDP Manipulation

Our main dependent variable is GDP manipulation, defined as the difference between reported GDP and true GDP. Data of reported GDP are retrieved from various yearbooks. Meanwhile, we recover the true GDP by utilizing the night light and air pollution satellite data, with a machine learning XGBoost technique. Night light and air pollution data are highly correlated with economic activity and can hardly be manipulated; thus, these data are widely used to predict and recover
true GDP records (Fowlie et al., 2019; Henderson et al., 2012; Morris and Zhang, 2018; Pinkovskiy et al., 2016). The method allows us to predict each county’s true GDP in the corresponding years and calculate the degree of GDP manipulation, respectively.\(^{10}\) Moreover, XGBoost outperforms traditional econometric methods because of its superior predictive power (Chen et al., 2019; Erel et al., 2018).\(^{11}\)

We also collect the personal information of officials as control variables, which include age, gender, years of education, tenure year, and counties’ socioeconomic characteristics such as population and GDP. We include the descriptive data summary in Table 1 below.

<table>
<thead>
<tr>
<th>Variables</th>
<th>Observations</th>
<th>Mean</th>
<th>SD</th>
<th>Min</th>
<th>Max</th>
</tr>
</thead>
<tbody>
<tr>
<td>GDP manipulation</td>
<td>5383</td>
<td>0.012</td>
<td>0.137</td>
<td>-0.380</td>
<td>0.598</td>
</tr>
<tr>
<td>Famine severity</td>
<td>5383</td>
<td>0.397</td>
<td>0.132</td>
<td>-0.009</td>
<td>0.862</td>
</tr>
<tr>
<td>Birth year</td>
<td>5383</td>
<td>1960</td>
<td>4.996</td>
<td>1942</td>
<td>1980</td>
</tr>
<tr>
<td>Famine exposure</td>
<td>5383</td>
<td>0.474</td>
<td>0.499</td>
<td>0</td>
<td>1</td>
</tr>
<tr>
<td>Birth place hit by famine</td>
<td>5383</td>
<td>0.469</td>
<td>0.499</td>
<td>0</td>
<td>1</td>
</tr>
<tr>
<td>Gender</td>
<td>5378</td>
<td>0.077</td>
<td>0.267</td>
<td>0</td>
<td>1</td>
</tr>
<tr>
<td>Education level</td>
<td>5080</td>
<td>5.428</td>
<td>0.716</td>
<td>2</td>
<td>8</td>
</tr>
<tr>
<td>Local origin</td>
<td>14892</td>
<td>0.141</td>
<td>0.348</td>
<td>0</td>
<td>1</td>
</tr>
<tr>
<td>Office Tenure</td>
<td>14892</td>
<td>2.377</td>
<td>1.541</td>
<td>1</td>
<td>14</td>
</tr>
</tbody>
</table>

### III. Results

#### A. Empirical Strategy

We use a difference-in-differences approach to identify the impact of famine exposure on officials’ GDP manipulative maneuvering. The approach utilizes the

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\(^{10}\) We justify the credibility of our estimates through three avenues. First, we show that the estimated manipulations are all positive when reported GDP are greater than the targets, and oscillate around zero otherwise. Second, our predictions are consistent with the revealed cases of manipulation. Third, our machine-learning-based estimates are also comparable to the night-light-based Henderson estimates. Illustrations of all the three justifications are included in the online appendix A.

\(^{11}\) It is a Scalable Tree Boosting System, which uses the sum of multiple tree regressions as predictions (Chen and Guestrin, 2016). We include the details of the practice in appendix A.
variations in the birth years and birthplaces of officials. First, by eliciting the differences in birth years, we compare and contrast officials based on whether they had childhood famine exposure. However, the simple comparison may elicit consequences beyond the famine, such as age-specific policy preferences. Therefore, we investigate birthplace variations, which calculate the differences between officials with and without childhood famine exposure in regions mildly hit by the famine and compare these with the differences between the two groups of officials in heavily-hit regions. The difference in such differences constitutes the causal effect of the famine. Specifically, our empirical strategy is summarized in the following equation, where $i$ represents an official, $c$ represents the birthplace, $y$ measures the degree of GDP manipulation in office, and $X$ is a series of controls. $\beta$ is the coefficient of interest, which captures the impact of famine exposure.

$$y_i = \alpha + \beta \text{Childhood}_i \times \text{Severity}_c + \gamma X_i + \epsilon_i$$  

(1)

B. Famine Exposure and GDP Manipulation: Baseline

We report our baseline results in Table 2. The first column presents the result with birth year and birthplace fixed effects. Indeed, we find a negative effect of famine exposure on GDP manipulation. In our sample horizon, officials with famine exposure are usually older. Therefore we add, in Column 2, the age in office to control age-specific risk preferences (per Greenwood and Nagel, 2009; Malmendier and Nagel, 2011). The negative impact of famine exposure remains robust. In Column 3, we further control gender and education level. In Column 4, we introduce provincial-level time trends to control for the unobservable time-varying factors at the provincial level. The negative impact of famine exposure remains robust in both specifications. Furthermore, mass political movements may reshape public beliefs and risk preferences (Costa-Font and Nicińska, 2019).
We thus include an interaction term between political radicalism\textsuperscript{12} and famine in Column 5, and show a limited correlation between political radicalism and GDP manipulation, while the impact of famine exposure remains significant. To cross-validate the credibility of our GDP manipulation estimates, we further adopt, in Column 6, the night light estimation to proxy GDP manipulation, as in Henderson et al. (2012). The result remains robust.

\textsuperscript{12} We follow Kung and Chen (2011) to use the count of “high-yield agricultural satellites” to proxy for political radicalism. The term is used commonly by mainstream Chinese media to describe agricultural production breakthroughs, which are often boasts of non-existent harvests reported by mainstream newspapers such as the People’s Daily (Ashton et al., 1984; Bernstein, 1984; Liang, 2003; Lu, 2008; Xie, 1990). We collect 658 records of such pseudo-achievements in the People’s Daily from August, 1958 to February, 1962.
Next, we explore the variations of the famine impacts across birth cohorts and regions. First, we break down the *Childhood* dummy into a continuous variable based on the year of birth and conduct a parallel trend test to estimate the year-by-year impact on each birth cohort. As shown in Figure 2, we find that famine had insignificant effects on cohorts born after 1962 since the entire cohort had no famine exposure. Meanwhile, the negative impacts for cohorts born before 1960 were converging to zero as birth years traced back since older officials tend to be more mature and, thus, had more stable and consistent risk preferences during the famine (Sahm, 2007; Stigler and Becker, 1977). Interestingly, there is a significant bounce from the 1961 cohort to the 1962 cohort. While it illustrates the impact of famine exposure vividly, it also suggests that even a famine exposure in a mother’s womb can change a child’s future behavior, echoing Currie and Vogl (2013).13

\[13\] This finding also echoes the literature on fetal origin hypothesis (Almond and Currie, 2011; Cronqvist et al., 2015). The fetal origin hypothesis refer to early life experiences such as nutrition conditions and diseases, may have long-run effects on health and well-being. Our finding extends the hypothesis to show that early life experiences may further reshape truth-telling behaviors.

**Table 2. Baseline Results**

<table>
<thead>
<tr>
<th></th>
<th>(1)</th>
<th>(2)</th>
<th>(3)</th>
<th>(4)</th>
<th>(5)</th>
<th>(6)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Manipulation</td>
<td>Manipulation</td>
<td>Manipulation</td>
<td>Manipulation</td>
<td>Manipulation</td>
<td>Night Light</td>
</tr>
<tr>
<td>Famine</td>
<td>-0.035 (0.011)</td>
<td>-0.034 (0.011)</td>
<td>-0.033 (0.011)</td>
<td>-0.040 (0.013)</td>
<td>-0.036 (0.019)</td>
<td>-0.013 (0.008)</td>
</tr>
<tr>
<td>Age</td>
<td>0.002 (0.001)</td>
<td>0.002 (0.001)</td>
<td>0.002 (0.001)</td>
<td>0.002 (0.001)</td>
<td>0.002 (0.001)</td>
<td>0.004 (0.001)</td>
</tr>
<tr>
<td>Female</td>
<td>-0.020 (0.013)</td>
<td>-0.021 (0.014)</td>
<td>-0.019 (0.014)</td>
<td>0.007 (0.010)</td>
<td>0.007 (0.010)</td>
<td>0.007 (0.010)</td>
</tr>
<tr>
<td>Education</td>
<td>0.003 (0.004)</td>
<td>0.002 (0.004)</td>
<td>0.002 (0.004)</td>
<td>-0.001 (0.003)</td>
<td>-0.001 (0.003)</td>
<td>-0.001 (0.003)</td>
</tr>
<tr>
<td>Famine*</td>
<td>-0.008 (0.022)</td>
<td>-0.008 (0.022)</td>
<td>-0.008 (0.022)</td>
<td>-0.008 (0.022)</td>
<td>-0.008 (0.022)</td>
<td>-0.008 (0.022)</td>
</tr>
<tr>
<td>Political Radicalism</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Observation</td>
<td>5,383</td>
<td>5,383</td>
<td>5,076</td>
<td>5,076</td>
<td>5,076</td>
<td>674</td>
</tr>
<tr>
<td>Birthplace FE</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
</tr>
<tr>
<td>Birthyear FE</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
</tr>
<tr>
<td>Time Trend</td>
<td>NO</td>
<td>NO</td>
<td>NO</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
</tr>
</tbody>
</table>
Relatedly, we extend the severity variable into the ten percentiles to investigate the heterogeneous effects of famine severity across regions. Figure B2 in the Online Appendix find that moderate famine exposures had limited impact on officials’ behavior, while severe famine exposures substantially restrained officials’ incentives to lie. Combining the temporal and regional variation analysis, we find that famine exposure restrained lying more effectively when such exposure is more severe, which is consistent with previous studies on disaster exposure (e.g., Kim and Lee, 2014).\textsuperscript{14}

\textbf{C. Robustness Checks}

We conduct four robustness checks and provide the details in the online appendix.

\textsuperscript{14} Relatedly, if the lying incentives are driven by career concerns, we expect officials to over-report GDP data. Consistent with the performance-based promotion schemes in China incentivize over-reports (Li and Zhou 2005, Xu 2011), we show that officials with famine exposure tend to restrain from over-, instead of under-reports (Appendix Table B1).
The first check speaks to an alternative explanation that individuals with certain pre-determined behavior, such as risk-tolerance, may manifest survival bias in the great famine and that the disaster selects risk-averse individuals to survive, instead of changing their behavior patterns.\textsuperscript{15} To rule out such selection effect, we narrow down our sample to include only the cohorts who were more than 6 years old when the famine hit. The narrowed sample features individuals with more stabilized behavior patterns such as patterns of risk preference and who are thus is more likely to be influenced by the selection effect other than a change of preferences. We find no significant results in the older cohorts, thus invalidating the selection effect (online appendix Figure B3).

The second check deals with prediction noise. A machine-learning-based algorithm may produce accurate predictions for samples close to the mean but and may feature notable errors for samples at extreme values. We thus use a concentrated sample which takes 200 and 250 sample points from the two sides of the cut-off, and replicate our baseline result. Our baseline result remains robust in the concentrated sample, which suggests that our results are not driven by prediction errors (online appendix Table B2).

The third check introduces additional controls on socioeconomic determinants to reduce the omitted variable bias. Our baseline result remains robust after controlling for local GDP, local population, whether the officials are new to the office, their tenure, and whether they are closer the years before retirement. We find that officials are more likely to lie in their first year in office (online appendix Table B3). This echoes consistently with the existing literature, as these are the officials with greater career concerns (Li and Zhou, 2005; Ferraz and Finan, 2011).

Lastly, we exercise a falsification test to confirm that our causal effect is not the consequence of spurious correlations due to noise or luck: by the random

\textsuperscript{15} The bias can be positive or negative. For example, risk-loving people may try various risky ways to get food and survive the famine, so those survivors can be more risk-loving. On the other hand, the ways that risk-loving people did may be illegal and they could get punished, so survivors can be risk-averse.
relocation of officials’ birthplace, we see no significant results of our coefficient of interest and confirm the non-significant results in 1,000, 3,000, 5,000, and 10,000 random trials (online appendix Figure B4).

IV. Mechanism

We believe the mechanism that prevents disaster-exposed politicians from lying is rooted in changes in individual risk preferences. Among existing studies, many have observed that disasters reshaped individuals to increase their aversion to risk and that such reshaping is persistent over time (van den Berg et al., 2009; Callen et al., 2014; Kim and Lee, 2014; Becchetti et al., 2017; Cassar et al., 2017; Brown et al., 2018; Guiso et al., 2018). Relatedly, a large psychology literature has pointed out that negative shocks in early life such as fear, rage, and helplessness may trigger an emotional excitement overload in the nervous system, leading to long-term risk aversion (Freud, 1957 [1915]; 1933; Hunt, 1979; Lerner et al., 2003). In addition to emotional responses to negative experiences, medical research also suggests negative shocks during childhood influence biological and genetic development, which in turn influences individuals’ risk preferences (Fox et al., 2010; Labonté et al., 2012; Mehta et al., 2013). Many lab experiments发现 truth-telling or lying aversion and risk-averse 的人更加可能 tell truth (Laine et al., 2013)。同时，中国严惩 GDP 造假：如果一个官员撒谎被 detect 到，那么他将受到严厉的惩罚，这意味着撒谎无疑是一个 risky 的行为。结合以上两点，Consequently, we conjecture that being exposed to one of the largest famines in human history may make the officials more risk-averse and, thus, more reluctant to lie.

As the preference profiles of the officials are private information, a direct test of the mechanism remains difficult. As a substitute, we employ the World Value Survey (WVS) data, which features a series of measures of individual preferences
and mindsets\textsuperscript{16} that lead to lying. In addition to risk preferences, other preferences include altruism, trust, social network, happiness, and time preferences. In particular, time preferences are measured by the patience to make long-term investments, such as expenditures on education, science and technology, public health, and social welfare. With the menu of preferences, we can not only investigate the impact of famine exposure but also compare it with those of other preferences using a similar design to our baseline analysis. To rule out the systematic differences between the WVS sample and our officials’ data, we compare several important covariates between the two datasets and control factors with systematic differences, such as educational levels.\textsuperscript{17}

Table 3 estimates the effects of famine experiences on individuals’ preferences and social and psychological factors, based on the WVS data. First, we provide confirmative evidence in Column 1 that famine exposure lowers individual risk preferences. Second, we find no evidence that other preferences such as altruism, trust, social network, or happiness are altered by the famine (Column 2-5). Consequently, our baseline results are not driven by these alternate preferences. To further address the potential systematic differences between officials and the general public, we conduct a propensity score weighting method (Hirano and Imbens, 2001; Hirano et al., 2003). The results are robust. In Columns 6-10, we show that the famine experiences did not change people’s long-term investment patterns and, thus, did not change their time preferences.\textsuperscript{18}

There is a concern that based on the WVS’s finding, the selection effect may confound. For this, we employed the method of Appendix Figure A3, which is based on the age group of those who are older than 6 years, to re-estimate Table 3. These individuals had more experience with disasters. Bellows and Miguel (2009), Cassar et al. (2011; 2013), Nock (1982) believe negative experience may influence trust to others. Time preference (Callen, 2015), social network (Belsky et al., 1991; Chen and Yang, 2018) and happiness (Adhvaryu et al., 2019) may also be influenced by disasters.\textsuperscript{17}

\textsuperscript{16} Specifically, Li et al. (2010) and Xu and Li (2016) find that people become more altruistic and cooperative after disasters. Bellows and Miguel (2009), Cassar et al. (2011; 2013), Nock (1982) believe negative experience may influence trust to others. Time preference (Callen, 2015), social network (Belsky et al., 1991; Chen and Yang, 2018) and happiness (Adhvaryu et al., 2019) may also be influenced by disasters.

\textsuperscript{17} The comparison can be found in online appendix Figure B5.

\textsuperscript{18} Existing literature also points out that disaster exposure may lead decision makers to be more ambiguous (Wakker, 2010, p. 354). In particular, a famine-exposed official may randomly select a GDP figure to report. We can rule out such ambiguity since our finding shows consistently the restrained manipulation of GDP reports.
selection effect。结果发现该子样本的风险偏好并没有被饥荒改变，这意味着 selection effect 非常有限。
To summarize, Table 3 provides supportive evidence that famine exposures rendered the exposed politicians more risk-averse, which in turn made them more reluctant to lie. Moreover, we find no evidence that the famine reshaped other social preferences, such as altruism, trust, social network, happiness, and time.
preferences.

V. Conclusion

The lies of public servants can be highly destructive to the public interest. To uncover the potential determinants of politicians’ lying incentives, this paper investigates the impact of famine exposure on GDP manipulation by county leaders in China. We show that leaders who experienced the great famine during 1959-61 showed greater restraint when manipulating local GDP data to achieve enhanced career advancement opportunities. Furthermore, by lending support from the WVS data, we show that exposure to the famine reshapes individuals to be more risk-averse, thus making them more reluctant to lie. The paper thus provides a first collection of quantitative evidence dealing with how disasters reshape the long-term well-being of the society through changing risk preferences that reduce the lying incentives of public servants.
Reference


Besley, Timothy, Torsten Persson, and Daniel M. Sturm. 2010. “Political


Cronqvist, Henrik, and Frank Yu. 2017. “Shaped by Their Daughters: Executives,


Mehta, D., T. Klengel, K. N. Conneely, A. K. Smith, A. Altmann, T. W. Pace, M. Rex-


Online Appendix A. Construction of GDP Manipulation

This section elaborates our method to calculate GDP manipulation. In short, we aim to recover the true GDP data from indicators that are closely related with economic growth (thus with GDP), and are less manipulable. In particular, we follow the estimation methods of Pinkovskiy and Sala-i-Martin (2016) and Clark et al. (2020). We start by collecting data of satellite light and air pollution – frequently used proxies for economic activities that are less manipulable by local officials. We include the sources of all data in appendix Table A1. We proceed to aggregate the data to county-level in annual bases, as the input variables of our GDP prediction. The input variables also include year and county dummies.

<table>
<thead>
<tr>
<th>Variables</th>
<th>Sources</th>
</tr>
</thead>
<tbody>
<tr>
<td>NO2 (3-year moving average)</td>
<td>European Space Agency (ESA)</td>
</tr>
<tr>
<td>NO2 (1-year moving average)</td>
<td>National Aeronautics and Space Administration (NASA)</td>
</tr>
<tr>
<td>SO2</td>
<td>European Commission (EC)</td>
</tr>
<tr>
<td>PM10</td>
<td>EC</td>
</tr>
<tr>
<td>PM2.5</td>
<td>NASA, EC</td>
</tr>
<tr>
<td>CO2</td>
<td>EC</td>
</tr>
<tr>
<td>OC</td>
<td>EC</td>
</tr>
<tr>
<td>BC</td>
<td>EC</td>
</tr>
<tr>
<td>CO</td>
<td>EC</td>
</tr>
<tr>
<td>NH3</td>
<td>EC</td>
</tr>
<tr>
<td>NMVOC</td>
<td>EC</td>
</tr>
<tr>
<td>NOx</td>
<td>EC</td>
</tr>
<tr>
<td>Population</td>
<td>Oak Ridge National Laboratory</td>
</tr>
<tr>
<td>Nighttime light</td>
<td>NASA</td>
</tr>
</tbody>
</table>

Next, we select the machine learning models. We divide the sample into 70% training set and 30% validation set, and use popular models such as random forest, support vector machine, neural network, and XGBoost for training. We then test the predictive power of each model, among which XGBoost outperforms the rest.
Therefore, we construct the outcome variables using XGBoost.

We justify the credibility of our estimates through three avenues. First, appendix Figure A1 depicts the distribution of GDP manipulation, where the horizontal axis marks GDP growth targets, and the vertical axis marks the frequency. As shown in the figure, the predicted manipulation is positive when the reported GDP is slightly above the target, and the gap shrinks when the reported GDP is far greater than the target. This is consistent with the intuition of manipulation, that individuals are strongly motivated to pass the threshold when their performance is slightly below the target, while the incentives of manipulation are less salient when the performance is well-above the target already. On the other hand, when the reported GDP is less than the target, the gap oscillates around zero – consistent with the presence of truth-telling officials.

![Figure A1. Distribution of GDP manipulations](image)

Note: We only present the gap to target GDP ratio between -10% and 10% for conciseness.

Next, we cross-compare our estimated manipulation with exposed GDP manipulation case in reality. As shown in appendix Figure A2, the red bars indicate the exposed provinces in reality – they are also predicted to have manipulated
local GDP in our estimation, which adds to the credibility of our estimates.

**Figure A2. GDP manipulations across provinces**

Lastly, we compare our calculated manipulations with the estimates by Henderson et al. (2012) based on satellite light data. The two estimates are highly correlated at 1% significance level.
Online Appendix B: Figures and Tables

Figure B1. GDP Distribution after Manual Revision of Targets

(A) Target GDP + 0.1%

(B) Target GDP – 0.1%
(C) Real target GDP

Notes: Panel (A) and (B) depict the distribution of the gap between reported GDP and the “placebo targets”: target GDP + 0.1% in Panel (A), and target GDP − 0.1% in Panel (B), respectively. Both distributions are smooth. However, Panel (C) depicts the distribution of the gap between reported GDP and the true targets: a jump exists around 0, which points to the possibility of manipulation.

Figure B2. Severity and Rectification

Notes: This figure replaces the binary famine exposure in the baseline to a generalized interaction term of the cohort dummies and deciles of famine severity, with the bottom 10% of the regions least hit by the famine as base group. As shown, when famine was not as severe, officials continued to lie in public offices. However, when the local famine was more severe, the officials became more
reluctant to lie when they served in public offices later on.

Figure B3. Selection Effects of the Famine

![Graph showing selection effects]

Notes: This figure re-estimate Figure B2 for samples in which individuals were more than six years old when the famine hit – a group with more stable risk preferences, thus the test captures the selection effects of the famine. As shown, the impacts are insignificant regardless of famine severity. Therefore, our baseline results are less likely driven by selection or survivor's bias.

Figure B4. Random Sample Trial

![Histograms for random sample trials]

1,000 times                                3,000 times

5,000 times                                10,000 times
Notes: The figure depicts the distribution of coefficients in random sample trials. Specifically, we randomly assign a famine severity to each county, and repeat our baseline according to the randomized famine. We replicate the random trials by 1,000, 3,000, 5,000, and 10,000 times, where Panel (A) to (D) present the distributions, respectively. Most of the coefficients are normally distributed, centring around zero, which suggests that our results are unlikely driven by spurious regression.

**Figure B5. Sample differences: Local officials vs. WVS**

Panel (A) compares the systematic differences of the officials’ cohort and the WVS respondents, in year of birth and educational levels. Such systematic differences urge us to use the propensity weighting method by Hirano and Imbens (2001) and Hirano et al. (2003) to address the issue in Table 3.
### Table B1. Direction of Manipulation

<table>
<thead>
<tr>
<th></th>
<th>(1) Absolute Value</th>
<th>(2) Over-reports</th>
<th>(3) Under-reports</th>
</tr>
</thead>
<tbody>
<tr>
<td>Childhood × Severity</td>
<td>-0.020 (0.009)</td>
<td>-0.039 (0.018)</td>
<td>-0.008 (0.015)</td>
</tr>
<tr>
<td>Observations</td>
<td>5,076</td>
<td>2,614</td>
<td>2,462</td>
</tr>
<tr>
<td>R-squared</td>
<td>0.348</td>
<td>0.498</td>
<td>0.525</td>
</tr>
<tr>
<td>Personal features</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
</tr>
<tr>
<td>Birth place Fixed Effects</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
</tr>
<tr>
<td>Birth year Fixed Effects</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
</tr>
<tr>
<td>Provincial Time Trend</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
</tr>
</tbody>
</table>

Notes: The econometric specification is similar with the baseline. In Column 1, the dependent variable is the absolute value of the manipulation, while Column 2 and 3 reports the results on the sub-sample of positive and negative manipulation, respectively. As shown, famine exposure mainly restrains manipulation driven by over-reports, instead of under-reports.

### Table B2. Concentrated Sample

<table>
<thead>
<tr>
<th></th>
<th>(1) 200 samples around cut-off</th>
<th>(2) 250 samples around cut-off</th>
</tr>
</thead>
<tbody>
<tr>
<td>Childhood famine exposure</td>
<td>-0.050 (0.027)</td>
<td>-0.045 (0.024)</td>
</tr>
<tr>
<td>Gender</td>
<td>-0.016 (0.011)</td>
<td>-0.015 (0.016)</td>
</tr>
<tr>
<td>Education</td>
<td>-0.006 (0.001)</td>
<td>-0.002 (0.004)</td>
</tr>
<tr>
<td>Constant</td>
<td>0.004 (0.011)</td>
<td>0.007 (0.051)</td>
</tr>
<tr>
<td>Observation</td>
<td>200</td>
<td>250</td>
</tr>
<tr>
<td>R-squared</td>
<td>0.601</td>
<td>0.578</td>
</tr>
<tr>
<td>Birth place Fixed Effects</td>
<td>YES</td>
<td>YES</td>
</tr>
<tr>
<td>Birth year Fixed Effects</td>
<td>YES</td>
<td>YES</td>
</tr>
<tr>
<td>Provincial Time Trend</td>
<td>YES</td>
<td>YES</td>
</tr>
</tbody>
</table>

Notes: This table re-estimates the baseline using two concentrated sub-samples, to rule out the impact of outliers: Column 1 with 200 samples around the cut-off where reported GDP equals target GDP; and Column 2 the other one with 250 samples. As shown, famine exposure restrains manipulation in both cases.
<table>
<thead>
<tr>
<th></th>
<th>(1)</th>
<th>(2)</th>
<th>(3)</th>
<th>(4)</th>
<th>(5)</th>
<th>(6)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Famine exposure</td>
<td>-0.025</td>
<td>-0.028</td>
<td>-0.023</td>
<td>-0.023</td>
<td>-0.020</td>
<td>-0.027</td>
</tr>
<tr>
<td></td>
<td>(0.010)</td>
<td>(0.010)</td>
<td>(0.012)</td>
<td>(0.013)</td>
<td>(0.010)</td>
<td>(0.010)</td>
</tr>
<tr>
<td>First year</td>
<td>0.005</td>
<td>0.003</td>
<td>0.003</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.003)</td>
<td>(0.002)</td>
<td>(0.002)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Years in office</td>
<td>-0.001</td>
<td>-0.002</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.001)</td>
<td>(0.001)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Retired</td>
<td></td>
<td></td>
<td></td>
<td>-0.031</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td>(0.019)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Observations</td>
<td>14,458</td>
<td>13,776</td>
<td>9,773</td>
<td>9,773</td>
<td>9,773</td>
<td>9,773</td>
</tr>
<tr>
<td>R-squared</td>
<td>0.400</td>
<td>0.405</td>
<td>0.491</td>
<td>0.491</td>
<td>0.491</td>
<td>0.491</td>
</tr>
<tr>
<td>FE’s</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
</tr>
<tr>
<td>Personal characteristics</td>
<td>NO</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
</tr>
<tr>
<td>Socioeconomic indicators</td>
<td>NO</td>
<td>NO</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
<td>YES</td>
</tr>
</tbody>
</table>

Notes: Column 1 presents the results without additional controls. Column 2-3 gradually add the individual characteristics (gender and educational levels) and socioeconomic indicators (logged local population and logged GDP per capita) as additional controls, while Column 4-6 add the features of officials’ career concerns. Column 4 adds a dummy taking value one if the official is in the first year of the tenure, and zero otherwise. This captures the extra incentives upon taking offices. Column 5 adds the number of years under which the official serves as the county party secretary, and Column 6 adds a dummy taking value one if the official is above 50 years old – and thus in the later stage of the career – and zero otherwise. The cut-off age of 50 years follows that in Yao and Zhang (2015). As shown, the effect of famine exposure remains robust in all columns.